Journal of Informatics and Mathematical Sciences

Vol. 10, No. 3, pp. 411–415, 2018 ISSN 0975-5748 (online); 0974-875X (print) Published by RGN Publications DOI: 10.26713/jims.v10i3.909



Research Article

# Stochastic Integrals and Power Contractions in Bernoulli Selections

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**Abstract.** Random contractions and Bernoulli selections are recognized as strong analytical tools of probability distributions theory. The paper investigates the distribution of a Bernoulli selection incorporating a stochastic integral and a random contraction. Moreover, the paper establishes a practical interpretation of the formulated Bernoulli selection.

Keywords. Stochastic integral; Random contraction; Bernoulli selection

MSC. 97K60; 97K50; 60E05

**Received:** May 7, 2018 **Accepted:** July 5, 2018

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# 1. Introduction

We consider the positive random variable H and the random variable V with values in the interval (0,1). We also consider the random variable

J = HV.

If the random variables H, V are independent then the random variable J is said a random contraction of the random variable H via the random variable V [7]. If the random variable V follows the power distribution then the random variable J is said a power contraction. Random contractions are generally recognized as strong analytical tools of probability theory for investigating unimodality [11], infinite divisibility [13], stability [5], selfedecomposability

[8] and other important properties of probability distributions. Moreover, random contractions have practical applications in income distributions analysis [10], cindynics [9], continuous discounting [4], reliability theory [3], inventory control [15], operations research [14], proactive risk management [1], engineering [12], systemics [16], and informatics [2].

The present paper is mainly devoted to the characterization of the distribution of a Bernoulli selection incorporating a random contraction and a stochastic integral.

## 2. Formulation of a Stochastic Model

The present section of the paper makes use of two positive random variables, a Bernoulli random variable and a stochastic integral in order to formulate a stochastic model.

Let  $\{X(t), t \ge 0\}$  be a stochastic process with stationary, independent, and positive increments. We assume that  $E(X(t)) = \mu t$  and  $V(X(t)) = \sigma^2 t$ . We also assume  $\{X(t), t \ge 0\}$  is continuous in probability and that its sample paths are right continuous and have left limits. Moreover, we assume that the increment L = X(t+1) - X(t) has characteristic function  $\varphi_L(u)$ . The stochastic integral

$$C = \int_0^\infty e^{-rt} dX(t), \quad r > 0$$

exists in the sense of convergence in probability and is finite almost surely. In addition, the distribution function of C is continuous and

 $\varphi_C(u) = \exp\left\{\int_0^\infty \log \varphi_L\left(ue^{-rt}\right)dt\right\}$ 

is its characteristic function [6]. The characteristic function  $\varphi_C(u)$  is easily shown to be given by

$$\varphi_C(u) = \exp\left\{\frac{1}{r}\int_0^u \frac{\log \varphi_L(w)}{w}dw\right\}.$$

Let N be a Bernoulli random variable with probability generating function

 $P_N(z) = q + pz, \quad 0$ 

S is a positive random variable with characteristic function  $\varphi_S(u)$  and T is a positive random variable with distribution function  $F_T(t)$ . We consider the random variable  $\Pi = Se^{-rT}$  and the stochastic model

$$Y = \begin{cases} \Pi, & N = 0, \\ C, & N = 1. \end{cases}$$

The following sections of the paper are devoted to the theoretical investigation and the practical interpretation of the formulated stochastic model.

## 3. Characteristic Function of a Stochastic Model

In general, an explicit evaluation of the characteristic function  $\varphi_Y(u)$  of the stochastic model Y is very difficult. The present section is devoted to the establishment of conditions for an explicit evaluation of a particular case of  $\varphi_Y(u)$ .

**Theorem.** We assume that the random variables N, C, S, T are independent and that  $F_T(t) = 1 - e^{-\lambda t}$ ,  $\lambda > 0$ . The characteristic function of the stochastic model Y has the form

$$\varphi_{Y}(u) = p \exp\left(\frac{1}{r} \int_{0}^{u} \frac{\log \varphi_{L}(w)}{w} dw\right) + q \left[\frac{ap}{u^{ap}} \int_{0}^{u} \exp\left(\frac{1}{r} \int_{0}^{w} \frac{\log \varphi_{L}(\theta)}{\theta} d\theta\right) w^{ap-1} dw\right]$$
with  $a = \lambda/r$  if and only if,  
 $Y \stackrel{d}{=} S$ ,
(1)

where  $\stackrel{d}{=}$  denotes equality in distribution.

*Proof.* Only the sufficiency condition will be proved since the necessity condition can be proved by reversing the argument. It is readily shown that the characteristic function of the stochastic model *Y* has the form

$$\varphi_Y(u) = p \exp\left(\frac{1}{r} \int_0^u \frac{\log \varphi_L(w)}{w} dw\right) + q \frac{a}{u^a} \int_0^u \varphi_S(w) w^{a-1} dw.$$
<sup>(1)</sup>

If we use (1) in (2) we get the integral equation

$$\varphi_Y(u) = p \exp\left(\frac{1}{r} \int_0^u \frac{\log \varphi_L(w)}{w} dw\right) + q \frac{a}{u^a} \int_0^u \varphi_Y(w) w^{a-1} dw.$$
(3)

If we multiply both sides of (3) by  $u^a$ ,  $u \neq 0$ , and then differentiating we get the differential equation

$$\varphi_{Y}(u) + \frac{u}{ap} \frac{d\varphi_{Y}(u)}{du} = \exp\left(\frac{1}{r} \int_{0}^{u} \frac{\log\varphi_{L}(w)}{w} dw\right) + \frac{u}{ar} \frac{\log\varphi_{L}(u)}{u} \exp\left(\frac{1}{r} \int_{0}^{u} \frac{\log\varphi_{Y}(w)}{w} dw\right).$$
(4)

From (4) we get that

$$\frac{ap}{u^{ap}} \int_{0}^{u} \varphi_{Y}(w) w^{ap-1} dw + \frac{ap}{u^{ap}} \int_{0}^{u} \frac{w}{ap} \frac{d\varphi_{Y}(w)}{dw} w^{ap-1} dw$$

$$= \frac{ap}{u^{ap}} \int_{0}^{u} \exp\left(\frac{1}{r} \int_{0}^{w} \frac{\log \varphi_{L}(\theta)}{\theta} d\theta\right) w^{ap-1} dw$$

$$+ \frac{ap}{u^{ap}} \int_{0}^{u} \frac{w \log \varphi_{L}(w)}{arw} \exp\left(\frac{1}{r} \int_{0}^{w} \frac{\log \varphi_{L}(\theta)}{\theta} d\theta\right) w^{ap-1} dw.$$
(5)

Moreover, it is readily shown that

$$\frac{ap}{u^{ap}} \int_0^u \frac{w}{ap} \frac{d\varphi_Y(w)}{dw} w^{ap-1} dw = \varphi_Y(u) - \frac{ap}{u^{ap}} \int_0^u \varphi_Y(w) w^{ap-1} dw$$
(6)

and that

$$\frac{ap}{u^{ap}} \int_{0}^{u} \frac{w \log \varphi_{L}(w)}{arw} \exp\left(\frac{1}{r} \int_{0}^{w} \frac{\log \varphi_{L}(\theta)}{\theta} d\theta\right) w^{ap-1} dw$$

$$= p \exp\left(\frac{1}{r} \int_{0}^{u} \frac{\log \varphi_{L}(w)}{w} dw\right) - \frac{ap^{2}}{u^{ap}} \int_{0}^{u} \exp\left(\frac{\log \varphi_{L}(\theta)}{\theta} d\theta\right) w^{ap-1} dw.$$
(7)
$$\exp\left(\frac{G}{u^{ap}}\right) \exp\left(\frac{1}{u^{ap}}\right) \exp\left($$

If we use (6) and (7) in (5) we get the characteristic function

$$\varphi_Y(u) = p \exp\left(\frac{1}{r} \int_0^u \frac{\log \varphi_L(w)}{w} dw\right) + q \left[\frac{ap}{u^{ap}} \int_0^u \exp\left(\frac{1}{r} \int_0^w \frac{\log \varphi_L(\theta)}{\theta} d\theta\right) w^{ap-1} dw\right]. \qquad \Box$$

It is obvious that the above characteristic function belongs to a Bernoulli selection incorporating the stochastic integral C and a power contraction of C.

# 4. Application

The theory of finance is concerned with the determination of the value of the firm as a going concern, the identification and analysis of factors with direct and indirect influence on this value, and with the valuation of investment opportunities. The economic value of the firm as a going concern is the present value of income that the firm will generate in the future. Assuming that the income is given by the stochastic process  $\{X(t), t \ge 0\}$  and since the corporate firm has an indefinite life, its economic value can be approximated by the stochastic integral C where r is the force of interest. Moreover, we assume that the random variable S denotes a cash flow arising at the random time T, then the random variable  $\Pi$  denotes the present value of S as viewed at time 0. Hence, the stochastic model formulated by the second section is suitable for making selection between two present values.

#### **Competing Interests**

The author declares that he has no competing interests.

#### **Authors' Contributions**

The author wrote, read and approved the final manuscript.

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